

Perspective

State Sepsis Mandates — A New Era for Regulation of Hospital Quality

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Sepsis is a major cause of illness and death in the United States, affecting more than 1.5 million Americans each year at an annual cost of over \$20 billion. To improve outcomes of sepsis,

policymakers are increasingly using regulatory mechanisms intended to provide incentives to clinicians and hospitals to improve the quality of sepsis care. One such initiative is an unprecedented set of New York State regulations implemented in 2013 and collectively known as "Rory's Regulations." Named after Rory Staunton, who died at 12 years of age from sepsis resulting from a soft-tissue infection, Rory's Regulations mandate that all hospitals in the state use evidence-based protocols for sepsis identification and management and that they report to the state government data on their sepsis-protocol adherence and clinical outcomes.

Rory's Regulations represent a major shift in the use of health policy to improve the quality of health care. Traditionally, governments seeking to improve health care quality have created marketbased incentives (such as public reporting of data on quality) or allocated resources for quality improvement (such as payments for "meaningful use" of health information technology). When governments are also the purchasers of health care, they can use pay-forperformance programs (such as Medicare's Value-Based Purchasing initiative) to encourage quality improvement. Rory's Regulations went a step further by legally mandating the use of specific guideline-based clinical practices — in this case, protocols that require timely fluid resuscitation, timely antibiotic administration, and frequent assessment of the hemodynamic response to therapy.¹

These regulations are still in their infancy, but a recent report issued by the New York State Department of Health provides some insight into their initial effects.² The report shows that the use of protocols for sepsis care in adults increased substantially from 2014 to 2016, from 73.7% to 84.7% of cases (see graphs). Compliance with early administration of intravenous fluids, early administration of antibiotics, and other elements of the early-resuscitation bundle increased from 41.5% to 55.2%, while mortality fell from 30.2% to 25.4%. Similar increases occurred in the use of care protocols in pediatric cases, although

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Rates of Initiation of Sepsis Protocol, Compliance with Protocol Bundle, and In-Hospital Mortality among Adult and Pediatric Hospitalized Patients with Sepsis in New York State, 2014–2016.

Data are from the New York State Department of Health.²

both compliance with the earlyresuscitation bundle and mortality remained relatively flat in the pediatric population.

New York Governor Andrew Cuomo lauded the results, stating in a press release that "our efforts are working to save lives and increase early detection and treatment of this deadly disease." Indeed, the report provides encouraging evidence that Rory's Regulations may be achieving their goal. However, several crucial questions remain, concerning not only the regulations' specific impact but also the broader question of the value of statewide mandates for protocolized sepsis care.

Although the report reveals that sepsis mortality in New York is decreasing, that decrease may not be attributable to the regulations. Epidemiologic data show that sepsis-related mortality has been decreasing throughout the United States, even in the absence of sepsis-focused regulation.³ Without data on sepsis-related mortality trends in New York prior to the regulations and comparable data from other states, it is impossible to say for certain that the observed changes in mortality are due to Rory's Regulations rather than being ongoing temporal trends.

Moreover, the potential adverse consequences of the regulations are still unknown. As Seymour et al. report in the Journal, the benefits of elements of protocolized early sepsis care such as blood culture, lactate measurement, and antibiotic administration are clear. However, recent clinical trials suggest that protocolized resuscitation strategies, which are also mandated by Rory's Regulations, may paradoxically lead to increased lengths of stay in the ICU and in the hospital and higher costs.4 The regulations may also lead to antibiotic overuse, if hospitals, in an attempt to increase their adherence to guidelines, give antibiotics to patients who are not infected. Antibiotic overuse may in turn increase the incidence of antibiotic

resistance and Clostridium difficile colitis, worsening clinical outcomes.

More broadly, there are insidious risks to turning clinical practice guidelines into policy mandates.⁵ Clinical practice guidelines may make strong recommendations based on rigorous scientific evidence, but they are also inherently flexible, allowing physicians to exercise considerable professional judgment.⁵ Legislation and regulation, however, are inherently inflexible, forcing clinicians to adopt certain care practices independent of clinical judgment. The best clinical practice guidelines originate from consensus among physicians and professional societies, whereas regulatory mandates may be politically motivated and may not incorporate the most recent scientific findings.⁵ To the degree that regulations encode practices that may later be shown to be nonbeneficial, or even harmful, they will subvert their purpose of improving health care quality.

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Another risk is that the regulatory process may be used to advance commercial interests that may not be in the public interest. Clinical practice guidelines often make recommendations involving proprietary medical devices and pharmaceuticals. Device and pharmaceutical companies could lobby state governments to include these products in future regulations. If these lobbying efforts are not transparent, conflicts of interest may lead to abuse of the regulatory process.5 Rory's Regulations were developed transparently with substantial input from New York hospitals and the clinical community. But the same may not be true in other states considering similar regulations in this or other domains of health care.

To address these issues, it is essential to rigorously evaluate the impact of Rory's Regulations in New York State, examining both their intended effects and unintended consequences. It is also essential to have a public debate regarding whether it is best to enforce clinical practice guidelines through regulation rather than allowing the medical community to decide whether to adopt them. In the meantime, it would make sense for states that adopt sepsis regulations to give hospitals maximal flexibility in deciding how to implement protocolguided sepsis care, to ensure that hospitals can respond appropriately as the evidence changes.

Sepsis is a public health crisis worthy of a policy response. Illinois, Pennsylvania, and Wisconsin are already in the process of issuing similar regulations, and the Rory Staunton Foundation, founded by the Staunton family, is actively seeking the implementation of sepsis-protocol mandates in every state by 2020. More direct efforts are needed to ensure that the government response to sepsis maximizes benefits, minimizes harms, and remains responsive to a complex and evolving evidence base.

Disclosure forms provided by the authors are available at NEJM.org.

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ORIGINAL ARTICLE

Time to Treatment and Mortality during Mandated Emergency Care for Sepsis

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ABSTRACT

BACKGROUND

In 2013, New York began requiring hospitals to follow protocols for the early identification and treatment of sepsis. However, there is controversy about whether more rapid treatment of sepsis improves outcomes in patients.

METHODS

We studied data from patients with sepsis and septic shock that were reported to the New York State Department of Health from April 1, 2014, to June 30, 2016. Patients had a sepsis protocol initiated within 6 hours after arrival in the emergency department and had all items in a 3-hour bundle of care for patients with sepsis (i.e., blood cultures, broad-spectrum antibiotic agents, and lactate measurement) completed within 12 hours. Multilevel models were used to assess the associations between the time until completion of the 3-hour bundle and risk-adjusted mortality. We also examined the times to the administration of antibiotics and to the completion of an initial bolus of intravenous fluid.

RESULTS

Among 49,331 patients at 149 hospitals, 40,696 (82.5%) had the 3-hour bundle completed within 3 hours. The median time to completion of the 3-hour bundle was 1.30 hours (interquartile range, 0.65 to 2.35), the median time to the administration of antibiotics was 0.95 hours (interquartile range, 0.35 to 1.95), and the median time to completion of the fluid bolus was 2.56 hours (interquartile range, 1.33 to 4.20). Among patients who had the 3-hour bundle completed within 12 hours, a longer time to the completion of the bundle was associated with higher risk-adjusted inhospital mortality (odds ratio, 1.04 per hour; 95% confidence interval [CI], 1.02 to 1.05; P<0.001), as was a longer time to the administration of antibiotics (odds ratio, 1.04 per hour; 95% CI, 1.03 to 1.06; P<0.001) but not a longer time to the completion of a bolus of intravenous fluids (odds ratio, 1.01 per hour; 95% CI, 0.99 to 1.02; P=0.21).

CONCLUSIONS

More rapid completion of a 3-hour bundle of sepsis care and rapid administration of antibiotics, but not rapid completion of an initial bolus of intravenous fluids, were associated with lower risk-adjusted in-hospital mortality. (Funded by the National Institutes of Health and others.)

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ORE THAN 1.5 MILLION CASES OF SEPsis occur in the United States annually, and many patients with sepsis present to the emergency department.¹ International clinical practice guidelines and the Centers for Medicare and Medicaid Services (CMS) recommend the prompt identification of sepsis and treatment with broad-spectrum antibiotic agents and intravenous fluids.^{2,3} These recommendations are supported by preclinical and observational studies suggesting that early treatment with antibiotics and intravenous fluids could reduce the number of avoidable deaths.^{4,5}

Yet, considerable controversy exists about how rapidly sepsis must be treated.⁶ Some clinicians question the potential benefit of rapid treatment, citing the absence of data from randomized trials, the potential for adverse effects, and the challenging implementation of these efforts in environments where staff are often overworked. Using data from New York,⁷ where hospitals are required to implement protocols and report on the treatment of sepsis, we examined the association between the timing of treatment and risk-adjusted mortality.

METHODS

STUDY DESIGN AND POPULATION

In early 2013, the New York State Department of Health (NYSDOH) began requiring hospitals to submit and follow evidence-informed protocols for the early identification and treatment of severe sepsis or septic shock (New York Codes, Rules, and Regulations parts 405.2 and 405.4). Although protocols could be tailored by each hospital, all the protocols were required to include a 3-hour bundle consisting of receipt of the following care within 3 hours: obtaining of a blood culture before the administration of antibiotics, measurement of the serum lactate level, and the administration of broad-spectrum antibiotics. Protocols were also required to include a 6-hour bundle, consisting of the administration of a bolus of 30 ml of intravenous fluids per kilogram of body weight in patients with hypotension or a serum lactate level of 4.0 mmol or more per liter, the initiation of vasopressors for refractory hypotension, and the remeasurement of the serum lactate level within 6 hours after the initiation of the protocol. Details about the treatment bundles are provided in Table S1 in the Supplementary Appendix, available with the full text of this article at NEJM.org.

We performed a retrospective study involving 185 hospitals in the NYSDOH database, including data from April 1, 2014, to June 30, 2016. All the hospitals were required to report patient-level data for patients with sepsis and septic shock to the Department of Health using electronic case-report forms that included data on demographic characteristics, coexisting conditions, characteristics of sepsis and septic shock, illness severity, and outcomes. Date and time stamps for protocol initiation and the elements of 3-hour and 6-hour bundled care were required for patients in whom a sepsis protocol was initiated. The state performed audits on a 10% random sample of hospitals using manual chart review and provided feedback to hospitals regarding data quality and completeness. Audit results are provided in Table S2 in the Supplementary Appendix. Patientlevel data were linked to hospital characteristics with the use of the NYSDOH administrative database. This study was approved with a waiver of informed consent by the NYSDOH institutional review boards.

SELECTION OF PATIENTS

Eligible encounters included those with patients who were older than 17 years of age and who had severe sepsis or septic shock, as defined with the use of criteria suggested in the 2001 International Sepsis Definitions Conference (Sepsis-2).8 In order to study only patients with communityacquired sepsis, we focused on patients who had a sepsis protocol initiated in the emergency department within 6 hours after arrival at the hospital. To remove outliers, we excluded patients in whom the 3-hour bundle was completed more than 12 hours after the initiation of the protocol. We also excluded patients in whom bundled care could be clinically contraindicated, patients with advance directives that limited treatment, patients who declined interventions, and patients who were enrolled in a clinical trial. We excluded 36 hospitals that had fewer than 50 cases of sepsis in order to remove spurious findings in reliability-adjusted models.9

Hospitals varied in their sepsis-identification strategies (see the Methods section in the Supplementary Appendix). These strategies included positive screening for sepsis on the basis of clinical assessment only (suspected or confirmed in-

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fection and two or more criteria for the systemic inflammatory response syndrome, with supporting laboratory test results optional); positive screening based on both clinical criteria and abnormal laboratory values; and a "code sepsis or rapid response" strategy that led to a positive screening based on clinical criteria. The regulations permitted hospitals to have flexibility with regard to case identification in order to facilitate broad adoption. Cases were not identified with the use of the Third International Consensus Definitions for Sepsis and Septic Shock (Sepsis-3) because these definitions were released after the implementation of the regulations was under way,¹⁰ and it was not possible to use post hoc adjudication. More than 98% of the patients with data entered in the database were confirmed to have had severe sepsis or septic shock on manual audit (Table S2 in the Supplementary Appendix). Cases that were found to have been entered erroneously could be removed by hospitals.

VARIABLES

The primary outcome was in-hospital mortality. The primary exposure was the time to completion of the 3-hour bundle, which was defined as the time in hours from the initiation of the protocol until all the elements of the 3-hour bundle were performed (i.e., blood cultures obtained, broad-spectrum antibiotics administered, and serum lactate level measured). If any element of the 3-hour bundle was performed before the start of the protocol, the patient was considered to have adhered to the protocol with regard to that element within the first hour. The time to the administration of broad-spectrum antibiotics was defined in a similar fashion. The time to the completion of the initial bolus of intravenous fluid was measured as the time from the initiation of the protocol until the completed administration of 30 ml of crystalloid fluid per kilogram, but only among patients who had a serum lactate level of 4.0 mmol or more per liter or who had hypotension (systolic blood pressure, <90 mm Hg).

Covariates included variables that were specified a priori as potential confounders between time to treatment and outcome on the basis of clinical experience and previous studies.^{10,11} These variables included demographic factors such as age, race or ethnic group, payer, burden of coexisting conditions, site of infection (e.g., respiratory, urinary, or skin), admission source (e.g., clinic, skilled nursing facility, or home), and measures of illness severity such as the presence of shock, serum lactate level, platelet count, or mechanical ventilation at admission. We developed a risk-adjustment model for in-hospital mortality using the above covariates with multivariable logistic regression that included a 90% random sample of the cohort. Internal validation of the model on the 10% remaining sample revealed adequate calibration (Hosmer–Lemeshow goodness-of-fit test with group size of 150, P=0.97) (Fig. S2 in the Supplementary Appendix) and discrimination (area under the receiver-operating-characteristic curve [C statistic], 0.77).

SENSITIVITY ANALYSES

We assessed the robustness of our analyses by repeating the primary analysis using the time to treatment as measured from the earliest recorded time of the presentation in the emergency department.6 We also assessed models in prespecified subgroups of patients. We repeated models with the subgroup of patients who had a protocol initiated up to 24 hours after arrival in the emergency department and with the subgroup of patients who had up to 24 hours between protocol initiation and completion of the 3-hour bundle.12 We repeated models with patients who were discharged to hospice care classified as dead at discharge and models that excluded any patients who had an element of the 3-hour bundle, administration of antibiotics, or completion of bolus of intravenous fluids before protocol initiation.

In supporting analyses, we measured the association of other elements of the 3-hour bundle with mortality, including the time to obtaining of a blood culture and the time to serum lactate measurement. We performed quantitative bias analysis to assess the magnitude of a hypothetical, unmeasured confounder that would be necessary to account for the association between the time to completion of the 3-hour bundle and risk-adjusted in-hospital mortality (see the Supplementary Appendix).¹³

STATISTICAL ANALYSIS

We performed bivariate analyses of the characteristics of the patients who had the 3-hour bundle in the emergency department completed within 3 hours and those who did not have the 3-hour bundle completed within that time win-

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dow. Continuous data are expressed as means with standard deviations or as medians with interquartile ranges, depending on normality. Categorical variables are shown as proportions. The range and variability in the times to treatments are shown with the use of histograms and cumulative proportions.

Multivariable modeling of the association between the time to treatment and in-hospital mortality was performed with the use of logistic regression, with adjustment for covariates. Binary variables were modeled as indicator covariates, and continuous variables were included as linear covariates, after assessment for nonlinear relationships with the use of fractional polynomials (P>0.05 for all models).¹⁴ We used multilevel regression with a random effect of hospital to account for hospital-level clustering. Each exposure (i.e., time to completion of the 3-hour bundle, time to the administration of broad-spectrum antibiotics, and time to completion of initial bolus of intravenous fluids) was evaluated separately. The risk of in-hospital death across the range of time to treatment was generated for the "typical" patient with the use of predictive margins that were adjusted for an average of the independent variables, as appropriate. We show adjusted risk estimates that are derived from the nonlinear models in order to show changes in risk over time.14

We used empirical Bayesian methods to determine the hospital-level rate of completion of the 3-hour bundle within 3 hours, administration of antibiotics within 3 hours, and completion of the initial bolus of intravenous fluids within 6 hours.⁹ We show the ranked order of adjusted rates across hospitals in caterpillar plots. All the analyses were performed with the use of Stata software, version 14.2 (StataCorp).

RESULTS

POPULATION OF PATIENTS AND TIME TO TREATMENT Of 111,816 patients at 185 hospitals, we excluded 21,046 patients (18.8%) who were ineligible, 32,665 (29.2%) who had protocols initiated outside the emergency department, 3648 (3.3%) who had protocols initiated after 6 hours, and 5126 (4.6%) who did not have the 3-hour bundle completed within 12 hours (Fig. S1 and Table S3 in the Supplementary Appendix). Of the remaining 49,331 eligible patients in the emergency department at 149 hospitals, most (40,696 patients [82.5%]) had the 3-hour bundle completed within 3 hours.





The 3-hour bundle for the care of patients with sepsis or septic shock had to include receipt of the following care within 3 hours: obtaining of a blood culture before the administration of antibiotics, measurement of the serum lactate level, and the administration of broadspectrum antibiotics; however, protocols could be tailored by each hospital. We also assessed the time to the administration of broad-spectrum antibiotics and the time to the completion of an initial bolus of intravenous fluids.

The median time to the completion of the 3-hour bundle was 1.30 hours (interquartile range, 0.65 to 2.35), the median time to the administration of broad-spectrum antibiotics was 0.95 hours (interquartile range, 0.35 to 1.95), and the median time to the completion of the initial bolus of intravenous fluids was 2.56 hours (interquartile range, 1.33 to 4.20) (Fig. 1). The characteristics of the patients who had the 3-hour bundle completed within 3 hours were similar to those who had the bundle completed during hours 3 through 12 (Table 1, and Table S4 in the Supplementary Appendix).

PRIMARY ANALYSES

In a multivariable model, each hour of time to the completion of the 3-hour bundle was associated with higher mortality (odds ratio of death until completion of 3-hour bundle, 1.04 per hour; 95% confidence interval [CI], 1.02 to 1.05; P<0.001) (Fig. 2, and Table S5 in the Supplementary Appendix). Patients who had the bundle completed during hours 3 through 12 had 14% higher odds

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Table 1. Characteristics of the Patients.				
Characteristic	All Patients (N=49,331)	3-Hr Bundle Completed in 3 Hr P Value		P Value*
		Yes (N=40,696)	No (N=8635)	
Percentage of patients	100.0	82.5	17.5	—
Age at admission — yr				<0.001
Median	73	73	71	
Interquartile range	60–83	61-84	59–82	
Female sex — no. (%)	23,634 (47.9)	19,157 (47.1)	4477 (51.8)	<0.001
Race — no. (%)†				<0.001
White	33,075 (67.0)	27,605 (67.8)	5470 (63.3)	
Black	8,269 (16.8)	6,487 (15.9)	1782 (20.6)	
Asian	2,167 (4.4)	1,774 (4.4)	393 (4.6)	
Other	5,820 (11.8)	4,830 (11.9)	990 (11.5)	
Hispanic ethnic group — no. (%)†	4,851 (9.8)	4,022 (9.9)	829 (9.6)	0.39
Coexisting condition — no. (%)				
Chronic respiratory failure	5,738 (11.6)	4,656 (11.4)	1082 (12.5)	0.004
Congestive heart failure	10,092 (<mark>20.5</mark>)	8,311 (20.4)	1781 (20.6)	0.67
End-stage renal disease	5,207 (10.6)	4,109 (10.1)	1098 (12.7)	<0.001
Admission source — no. (%)				<0.001
Home	33,464 (67.8)	27,306 (67.1)	6158 (71.3)	
Skilled nursing facility	13,233 (<mark>26.8</mark>)	11,247 (27.6)	1986 (23.0)	
Other‡	2,634 (5.3)	2,143 (5.3)	491 (5.7)	
Site of infection — no. (%)				<0.001
Urinary	13,439 (<mark>27.2</mark>)	10,963 (26.9)	2476 (28.7)	
Respiratory	19,839 (<mark>40.2</mark>)	16,806 (41.3)	3033 (35.1)	
Gastrointestinal	4,649 (9.4)	3,580 (8.8)	1069 (12.4)	
Other∬	11,404 (23.1)	9,347 (23.0)	2057 (23.8)	
Positive blood cultures — no. (%)	14,574 (29.5)	12,322 (30.3)	2252 (26.1)	<0.001
Serum lactate — mmol/liter				<0.001
Median	2.7	2.8	2.5	
Interquartile range	1.7-4.4	1.8-4.4	1.6-4.1	
Septic shock — no. (<mark>%</mark>)	22,336 (45.3)	18,393 (45.2)	3943 (45.7)	0.43
Teaching facility — no. (%)	40,257 (81.6)	7,739 (19.0)	7300 (84.5)	<0.001
In-hospital death — no. (%)	11,251 (22.8)	9,213 (22.6)	2038 (23.6)	0.05

* P values are based on Pearson's chi-square test for categorical variables and the Wilcoxon rank-sum test for continuous variables.

† Race and ethnic group were determined from medical records.

‡ Other locations include clinic or unknown.

J Other site of infection includes skin, central nervous system, and unknown.

of in-hospital death than patients in whom all mortality was similar (odds ratio of death until three items in the 3-hour bundle were completed antibiotics were administered, 1.04 per hour; in 3 hours (odds ratio, 1.14; 95% CI, 1.07 to 1.21; 95% CI, 1.03 to 1.06; P<0.001) (Fig. S3 in the P<0.001). The association between the time to Supplementary Appendix). Patients who received the administration of antibiotics and in-hospital antibiotics in hours 3 through 12 had 14% higher

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Subgroup	No. of Patients	Odds Ratio (95%)	CI)		
All patients	49,331	⊢ ●-1	1.04 (1.02–1.05)		
Sex			· · · ·		
Male	25,689	⊢● →	1.04 (1.02–1.05)		
Female	23,634	⊢ ●i	1.03 (1.02-1.05)		
Vasopressor use					
Yes	16,721	-	1.05 (1.03-1.07)		
No	32,610	↓ ●-1	1.02 (1.00-1.03)		
Admission source					
Home	33,464	⊢ ●i	1.04 (1.02-1.05)		
Other	15,867	⊢ ●i	1.04 (1.02–1.06)		
Coexisting condition					
Congestive heart failure	10,092	⊢ •−−+	1.06 (1.04-1.09)		
Hemodialysis	5,207	⊢	1.06 (1.03-1.09)		
Chronic respiratory failure	5,738	⊢	1.06 (1.03-1.09)		
Site of infection					
Respiratory	19,839		1.03 (1.01-1.06)		
Urinary	13,439	⊢	1.03 (1.01-1.06)		
Other	16,053	⊢● →	1.04 (1.02-1.06)		
Bacteremia					
Gram positive	7,175	⊢↓ ●	1.01 (0.98-1.05)		
Gram <mark>negative</mark>	6,431	⊢	1.05 (1.01-1.09)		
Other	965	·•	1.15 (1.07–1.24)		
None	34,757	⊢ ●-i	1.03 (1.02-1.05)		
		1.0 1.1 1.3			
			→		
	In-Hospita	I Death In-Hospital Death			
	Less Li	Kely More Likely			
Figure 2. Risk-Adjusted Odds Ratios of In-Hospital Death in the Primary Model and Prespecified Subgroups.					
Shown are odde ratios, with 05% confidence intervals, for in hospital death for each how that it took to complete					
shown are outs ratios, with 55% confidence mervars, to in-hospital death for each four that it took to complete					

odds of in-hospital death than those who received antibiotics within 3 hours (odds ratio, 1.14; 95% CI, 1.06 to 1.22; P=0.001). These associations appeared to be stronger among patients receiving vasopressors than among those who were not receiving vasopressors (Fig. 2, and Fig. S3 in the Supplementary Appendix). Figure 3 shows the crude and predicted risks of in-hospital death across a range of times to treatment in typical patients who presented to the emergency department. On average, the completion of the 3-hour bundle at 6 hours was associated with mortality that was approximately 3 percentage points higher than the mortality associated with completion of the bundle within the first hour.

Among the 26,978 patients who were eligible for and had the bolus of intravenous fluids completed within 12 hours, the time to completion of the fluid bolus was not associated with inhospital mortality (odds ratio of death until fluid bolus was completed, 1.01 per hour; 95% CI, 0.99 to 1.02; P=0.21) (Fig. S4 in the Supplementary Appendix). Patients who had the initial fluid bolus completed during hours 6 through 12 had an odds of in-hospital death that was similar to that among patients who had the initial fluid bolus completed within 6 hours (odds ratio of death for >6 hours to complete intravenous-fluid bolus, 1.02; 95% CI, 0.92 to 1.14; P=0.65). We found no interaction between time to the administration of antibiotics and time to completion of the initial bolus of intravenous fluids (P=0.88).

ADDITIONAL ANALYSES

A sensitivity analysis that used the earliest time of arrival in the emergency department to measure the time to treatment showed an association that was similar to that in the primary analyses. The results were unchanged when hospice discharges were reclassified as in-hospital deaths or when we excluded patients who had treatments completed before protocol initiation. When the

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time window for protocol initiation or completion of the 3-hour bundle was relaxed to 24 hours, the association between completion of the bolus of intravenous fluids and mortality became significant, albeit of very small magnitude (odds ratio 1.001; 95% CI, 1.000 to 1.002; P=0.03). Details are provided in Table S6 in the Supplementary Appendix.

In supporting analyses, we found that the time to obtaining a blood culture was associated with mortality (odds ratio, 1.04 per hour; 95% confidence interval, 1.02 to 1.06; P<0.001). Similar findings were observed for each hour until serum lactate measurement (Figs. S5 and S6 in the Supplementary Appendix). The quantitative bias analysis indicated that our results would be robust unless an unmeasured confounder was at least twice as prevalent among patients who had the 3-hour bundle completed later as among those who had it completed 1 hour earlier and unless the unmeasured confounder increased the odds of in-hospital death by more than 1.35 times (Fig. S7 in the Supplementary Appendix).

The risk-adjusted and reliability-adjusted rates of completing the 3-hour bundle ranged from 53 to 97% (median, 83%; interquartile range, 75 to 88) (Fig. 4, and Fig. S8 in the Supplementary Appendix). After we ranked hospitals from the lowest to greatest likelihood of completing the 3-hour bundle, the hospitals in the highest decile, despite similar illness severity among their patients, were 1.5 times as likely to complete the 3-hour bundle as hospitals in the lowest decile (94.3% vs. 64.1%). Hospitals that had a higher rate of bundle completion within 3 hours were somewhat smaller and less likely to be teaching hospitals than those that took longer than 3 hours to complete the bundle (Table S7 in the Supplementary Appendix).

DISCUSSION

Our findings support an association between time to treatment and outcome among patients with sepsis or septic shock treated in the emergency department during a statewide initiative mandating protocolized care. We found that a longer time to completion of a 3-hour bundle of care for patients with sepsis and the administration of broad-spectrum antibiotics were each associated with higher risk-adjusted in-hospital mortality. In our primary analysis, we did not find an association between the time to completion



Figure 3. Crude In-Hospital Mortality and Predicted Risks of In-Hospital Death.

Shown are the crude in-hospital mortality and predicted risks of in-hospital death, with adjustment for covariates across a range of time after protocol initiation, for the completion of the 3-hour bundle of sepsis care (Panel A), the administration of broad-spectrum antibiotics (Panel B), and the completion of the initial bolus of intravenous fluids (Panel C) in a typical patient. I bars represent 95% confidence intervals.

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Figure 4. Reliability-Adjusted Rate for Each Hospital for Completion of the 3-Hour Bundle in 3 Hours, According to Hospital Rank.

The 149 hospitals that were included in the study were ranked from lowest to highest, with higher numbers indicating a greater likelihood of completing the 3-hour bundle within 3 hours. I bars represent 95% confidence intervals.

of the initial **bolus** of intravenous **fluids** and inhospital mortality. The time to treatment varied widely across hospitals.

Our findings are **consistent** with multiple smaller, observational studies.^{5,15,16} A recent meta-analysis of 11 observational studies, however, showed no significant mortality benefit of the administration of antibiotics within 3 hours, as compared with after 3 hours, after triage in the emergency department (odds ratio, 1.16; 95% CI, 0.92 to 1.46) or within 1 hour after the recognition of shock (odds ratio, 1.46; 95% CI, 0.89 to 2.40).⁶ The odds ratios we report are similar, but the confidence intervals are narrower given the much larger sample size that was included in our study.

This study complements a patient-level metaanalysis of goal-directed therapy in severe sepsis and septic shock, the Protocolized Resuscitation in Sepsis Meta-Analysis (PRISM) trial.¹⁷ More than three of four patients in the PRISM trial received elements of the 3-hour bundle before randomization, after which the various trials composing the PRISM trial tested whether protocolized resuscitation strategies improved outcomes. Our study asked a different question: does timing matter for these earliest and most basic elements of care? These population-level data also place in context the relatively high compliance with these steps in the control groups of the various trials composing the PRISM trial before randomization. Only half the hospitals in the statewide database performed near this level.

There are several biologic explanations for the association between the time to completion of a 3-hour treatment bundle and outcome. First, more rapid administration of antibiotics reduces pathogen burden, modifies the host response, and could reduce the incidence of subsequent organ dysfunction. Second, clinicians who decide more quickly to measure the serum lactate level may identify heretofore unrecognized shock and are more prepared to deliver lactate-guided resuscitation than clinicians who are slower to measure the serum lactate level — a strategy that may improve outcome in randomized trials.¹⁸ Third, physicians have broad variation in how they identify sepsis, even when they are presented with similar cases.¹⁹ Fast delivery of sepsis treatment, even within the structure of mandated protocols, requires a prompt clinical suspicion of both infection and worsening organ dysfunction.

Although we found no association between the time to completion of the initial bolus of intravenous fluids and outcome in our primary analysis, these data should not be interpreted as evidence in favor of abandoning early fluid resuscitation. The analysis of the time to completion of the initial fluid bolus is most prone to confounding by indication (e.g., sicker patients will receive fluids sooner and are also more likely to die).²⁰ A greater volume of fluids given at rapid pace may also contribute to adverse effects such as pulmonary edema, volume overload, and longer duration of organ support in selected patients.²¹ Causal inference will require investigation in randomized clinical trials, and our analysis contributes to the clinical equipoise needed for such trials.

We found a variation of 1 to 2 times across hospitals with regard to the rates of completing the 3-hour bundle, the administration of antibiotics, and the completion of a bolus of intravenous fluids in the emergency department. Adherence, in general, ranged from 60 to 90%, and was greater than in comparable quality-improvement

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programs for stroke treatment in New York.²² Such performance may stem from increasing public awareness and advocacy about sepsis and national quality-improvement initiatives led by CMS.²³ Adherence was greatest in the emergency departments at smaller nonteaching hospitals, a finding that differs from a previous cohort study.²⁴ These hospitals may have fewer clinicians to train, a lower census in the emergency department, and a different case mix as compared with larger referral centers, which perhaps facilitates the more rapid implementation of sepsis protocols.

Our study has several limitations. First, this was not a randomized trial, so the results may be biased by confounding. Of greatest concern may be the lack of data about the appropriateness of broad-spectrum antibiotics. The appropriateness of the initial choice of an antibiotic agent has been associated with risk-adjusted mortality²⁵ but may be measurable only in the minority of patients with positive cultures and may differ according to local pathogen and antimicrobial resistance profiles. The hospitals included in this study were limited to a single state that may have epidemiologic features of sepsis that are distinct from those in other geographic regions.²⁶ The start time for measuring delays may not be accurate in all cases.

To address this, we evaluated models that used the earliest time of arrival in the emergency department and found no change in associations.

Our statewide evaluation showed that the times to the completion of a 3-hour bundle and the administration of broad-spectrum antibiotics were associated with greater in-hospital mortality among patients with severe sepsis and septic shock in the emergency department. We found no association between the time to completion of the initial bolus of intravenous fluids and outcome. If the relationship is causal, prompt recognition and faster treatment of sepsis and septic shock in the context of emergency care may reduce the incidence of avoidable deaths.

The views expressed in this article are those of the authors and do not necessarily represent the view of the U.S. government or the Department of Veterans Affairs.

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