# EDITORIAL



# **Transfusion Thresholds in FOCUS**

Paul J. Barr, Ph.D., and Karen E.M. Bailie, M.D., Ph.D.

Even though red-cell transfusion is an accepted and widely used intervention, questions regarding which patients should receive transfusions and under what circumstances continue to spark debate. There has been limited evidence from clinical trials to inform policy. Meanwhile, concern about potential risks of blood transfusion and the costs of maintaining an adequate and safe blood supply<sup>1</sup> have heightened interest in strategies to reduce the use of red-cell transfusion. These strategies include preoperative optimization of hemoglobin levels, the use of cell salvage during and after surgery, and pharmacologic interventions to reduce blood loss.

In the Transfusion Trigger Trial for Functional Outcomes in Cardiovascular Patients Undergoing Surgical Hip Fracture Repair (FOCUS),<sup>2</sup> Carson et al. describe another strategy: the use of lower hemoglobin thresholds to initiate transfusion. A recent Cochrane review1 of 17 trials comparing variously defined transfusion triggers, including a lower hemoglobin threshold (range, 7 to 9 g per deciliter) and a higher hemoglobin threshold (range, 9 to 12 g per deciliter), concluded that although there was good evidence that lower thresholds reduced blood use, available data were inadequate to determine whether lower thresholds had adverse effects on functional status or other major outcomes.

Although the review suggested that it was likely that a restrictive hemoglobin threshold for transfusion was as safe as a liberal strategy, the results were heavily <u>weighted</u> by the Transfusion Requirements in Critical Care <u>(TRICC)</u> trial.<sup>3</sup> The TRICC trial compared a threshold of <u>10</u> g per deciliter with a threshold of 7 g per deciliter in patients in the intensive care unit (ICU). Overall, 30-day rates of death were similar in the two groups, but in the predefined subgroups of patients <u>under the age of 55</u> years and those who were less critically ill, rates of death were significantly lower in the restrictive-strategy group. Complications occurring in the ICU were also similar overall, with the notable exception of a significantly higher number of cardiac events in the <u>restrictive-strategy</u> group. Thus, the overall results suggesting similar outcomes with a lower versus higher transfusion threshold might <u>not</u> be broadly <u>generalizable beyond</u> the ICU setting and particularly to patients with <u>cardiac</u> disease.

The FOCUS trial was designed to address this question among patients who had a history of or risk factors for <u>ischemic</u> heart disease and were undergoing surgical repair of a <u>fractured</u> <u>hip.<sup>2,4</sup></u> This study population is an important one in which to pose this question, since such patients are likely to be transfused because of surgical blood loss and their age, and it is plausible that they may be compromised at lower hemoglobin levels.<sup>5-7</sup>

In this study, 2016 patients over the age of 50 years (mean, 81.6) were randomly assigned to two study groups once their postoperative hemoglobin level fell below 10 g per deciliter. In the liberal-strategy group, single-unit transfusions were given to restore and maintain a hemoglobin level above 10 g per deciliter. In the restrictivestrategy group, transfusions were given when the hemoglobin level fell below 8 g per deciliter. Transfusion for symptoms of anemia was permitted in both groups. The trial had good statistical power to detect a difference in the primary outcome (death or an inability to walk 10 ft

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without human assistance) at <u>60 days</u> after randomization, which was ascertained through telephone calls by assessors who were unaware of study-group assignments in all but 17 participants.

Although the restrictive-strategy group received only half the number of transfusions administered in the liberal-strategy group, rates of the primary outcome did not differ significantly between the two groups, with 35% of patients in each group unable to walk unassisted, including 142 patients (76 of whom were in the liberalstrategy group) who had died. However, the average hemoglobin level in both groups was low by World Health Organization standards8 and could have limited functional recovery regardless of the transfusion strategy.6,7 There were no significant differences between groups among predefined secondary and tertiary outcomes, including inhospital myocardial events, other coexisting illnesses, and final discharge destination, but the study was not adequately powered to assess these outcomes.

Although a significant between-group separation in hemoglobin levels was obtained, the <u>absolute difference</u> was on average <u>only 1 g</u> per deciliter. This difference may not have been clinically significant in this group of elderly patients and therefore may have contributed to the lack of detectable difference in the primary outcome between the two groups.

Although the interpretation that the transfusion strategy makes no difference in functional outcome in patients without symptoms of anemia is likely to be correct, the same cannot be said with confidence of the lack of adverse effects, given the study's insufficient power to assess these events. Nevertheless, the absolute numbers of adverse events in this high-risk group was relatively low; cardiovascular events, including myocardial ischemia, heart failure, transient ischemic attack, and stroke, each occurred in no more than 6% of patients in either group. This low event rate is reassuring and supports the authors' conclusion that a postoperative hemoglobin threshold of 8 g per deciliter in the absence of symptomatic anemia appears to be acceptable in elderly patients with or at risk of ischemic heart disease.

In using lower hemoglobin thresholds to guide transfusion, the risks of undertransfusion should not be overlooked. The decision to transfuse should be guided by an assessment of individual patients on the basis of a combination of signs, symptoms, and laboratory measures, and not by a single hemoglobin level. In the FOCUS trial, transfusion for symptoms occurred more often in the restrictive-strategy group, and protocol violations resulting in additional transfusions (a total of 56) in this group may have reduced this frequency.

Despite these caveats, the FOCUS trial provides new evidence to support the view that a more restrictive transfusion threshold in the absence of symptoms of anemia may be reasonable, including in elderly patients at risk for cardiovascular events. Such a policy would reduce exposure to allogeneic blood transfusion with its attendant risks and costs.

Disclosure forms provided by the author are available with the full text of this article at NEJM.org.

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# ORIGINAL ARTICLE

# Liberal or Restrictive Transfusion in High-Risk Patients after Hip Surgery

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ABSTRACT

#### BACKGROUND

The hemoglobin threshold at which postoperative red-cell transfusion is warranted is controversial. We conducted a randomized trial to determine whether a higher threshold for blood transfusion would improve recovery in patients who had undergone surgery for hip fracture.

#### METHODS

We enrolled 2016 patients who were 50 years of age or older, who had either a history of or risk factors for cardiovascular disease, and whose hemoglobin level was below 10 g per deciliter after hip-fracture surgery. We randomly assigned patients to a liberal transfusion strategy (a hemoglobin threshold of 10 g per deciliter) or a restrictive transfusion strategy (symptoms of anemia or at physician discretion for a hemoglobin level of <8 g per deciliter). The primary outcome was death or an inability to walk across a room without human assistance on 60-day follow-up.

#### RESULTS

A median of 2 units of red cells were transfused in the liberal-strategy group and none in the restrictive-strategy group. The rates of the primary outcome were 35.2% in the liberal-strategy group and 34.7% in the restrictive-strategy group (odds ratio in the liberal-strategy group, 1.01; 95% confidence interval [CI], 0.84 to 1.22), for an absolute risk difference of 0.5 percentage points (95% CI, -3.7 to 4.7). The rates of in-hospital acute coronary syndrome or death were 4.3% and 5.2%, respectively (absolute risk difference, -0.9%; 99% CI, -3.3 to 1.6), and rates of death on 60-day follow-up were 7.6% and 6.6%, respectively (absolute risk difference, 1.0%; 99% CI, -1.9 to 4.0). The rates of other complications were similar in the two groups.

#### CONCLUSIONS

A liberal transfusion strategy, as compared with a restrictive strategy, did not reduce rates of death or inability to walk independently on 60-day follow-up or reduce in-hospital morbidity in elderly patients at high cardiovascular risk. (Funded by the National Heart, Lung, and Blood Institute; FOCUS ClinicalTrials.gov number, NCT00071032.)

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N THE UNITED STATES, MORE THAN 17 MILlion red-cell units are collected annually, and 15 million units are transfused.<sup>1</sup> Blood transfusions are frequently given to surgical patients and to the elderly.<sup>2,3</sup> Yet, the indications for postoperative transfusion have not been adequately evaluated and remain controversial. Most clinical trials have been small.<sup>4</sup> One adequately powered trial involving adults in intensive care units showed a nonsignificant decrease in 30-day mortality with a restrictive transfusion strategy, as compared with a liberal strategy (18.7% vs. 23.3%).5 However, the effect of a restrictive approach on functional recovery or risk of myocardial infarction in patients with cardiac disease has not been studied.4 We performed the Transfusion Trigger Trial for Functional Outcomes in Cardiovascular Patients Undergoing Surgical Hip Fracture Repair (FOCUS) to test the hypothesis that a higher threshold for blood transfusion (a hemoglobin level of 10 g per deciliter) would improve functional recovery and reduce morbidity and mortality, as compared with a more restrictive transfusion strategy (a hemoglobin level of <8 g per deciliter or symptoms).

#### METHODS

#### PATIENTS

From July 19, 2004, through February 28, 2009, we enrolled patients at 47 clinical sites in the United States and Canada. Telephone follow-up ended on May 4, 2009. Patients 50 years of age or older who were undergoing primary surgical repair of a hip fracture and who had clinical evidence of or risk factors for cardiovascular disease were eligible if they had a hemoglobin level of less than 10 g per deciliter within 3 days after surgery. According to the original protocol, only patients with cardiovascular disease (a history of ischemic heart disease, electrocardiographic evidence of previous myocardial infarction, a history or presence of congestive heart failure or peripheral vascular disease, or a history of stroke or transient ischemic attack) were eligible. In December 2005, eligibility criteria were expanded to enhance recruitment by including patients with any of the following cardiovascular risk factors: a history of or treatment for hypertension, diabetes mellitus, or hypercholesterolemia; a cholesterol level of 200 mg or more per deciliter or a lowdensity lipoprotein cholesterol level of 130 mg or

more per deciliter; current tobacco use; or a creatinine level of more than 2.0 mg per deciliter.<sup>6</sup>

We excluded patients if they were unable to walk without human assistance before hip fracture, declined blood transfusions, had multiple trauma (defined as having had or planning to undergo surgery for non-hip-related traumatic injury), had a pathologic hip fracture associated with cancer, had a history of clinically recognized acute myocardial infarction within 30 days before randomization, had previously participated in the trial with a contralateral hip fracture, had symptoms associated with anemia (e.g., ischemic chest pain), or were actively bleeding at the time of potential randomization.

The institutional review board or ethics committee at all 47 participating clinical sites approved the protocol (available with the full text of this article at NEJM.org). An independent data and safety monitoring board also approved the protocol and monitored the trial. Written informed consent was obtained from patients or their designated representatives. Methods were reported in detail previously.<sup>6</sup>

#### TREATMENT ASSIGNMENT AND FOLLOW-UP

We randomly assigned patients to the liberalstrategy group or the restrictive-strategy group using an automated telephone randomization system. Staff members at the data coordinating center prepared randomization schedules for each site using randomly ordered block sizes of two, four, six, or eight. After randomization, clinical-site staff members, clinicians, and patients were aware of study-group assignments.

Patients in the liberal-strategy group received 1 unit of packed red cells and additional blood as needed to maintain a hemoglobin level of 10 g or more per deciliter. An assessment of the hemoglobin level after transfusion was required, and an additional unit of blood was transfused if the patient's hemoglobin level was below 10 g per deciliter.

Patients in the restrictive-strategy group were permitted to receive transfusions if symptoms or signs of anemia developed or at the discretion of their physicians if the hemoglobin level fell below 8 g per deciliter. Symptoms or signs that were considered indications for transfusion were chest pain that was deemed to be cardiac in origin, congestive heart failure, and unexplained tachy-

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cardia or hypotension unresponsive to fluid replacement. Blood was administered 1 unit at a time, and the presence of symptoms or signs was reassessed. Patients with clinically diagnosed dementia received transfusions when the hemoglobin level fell below 8 g per deciliter because they might not be able to report their symptoms.

Hemoglobin levels were measured during hospitalization on days 1, 2, 4, and 7 after randomization. Additional hemoglobin determinations were made as clinically indicated. The assigned transfusion strategy was to be followed until discharge or up to 30 days, whichever came first. Transfusion was permitted at any time without measuring a hemoglobin level if the patient was bleeding and emergency transfusion was considered necessary by the treating physician.

Nurses at the clinical coordinating center who were not involved with study implementation and were unaware of study-group assignments telephoned patients or proxies at or close to 30 days and 60 days after randomization to ascertain outcomes after hospital discharge. They spoke directly to patients who were accessible by telephone or to proxies if patients were cognitively impaired or could not talk on the telephone.

#### PRIMARY OUTCOME

The primary outcome was death or an inability to walk 10 ft (or across a room) without human assistance at the 60-day follow-up. We hypothesized that an increased hemoglobin level would allow patients to participate more actively in rehabilitation and therefore increase the proportion who were walking independently 60 days after randomization.

#### SECONDARY OUTCOMES

Secondary outcomes included a combined outcome of in-hospital myocardial infarction, unstable angina, or death for any reason; each of these outcomes was assessed individually.

Electrocardiography was performed before surgery, before randomization, and on day 4 after randomization (or at the time of discharge if before day 4). Blood (plasma or serum) specimens were collected for measurement of the cardiac troponin I level before surgery, before randomization, and on days 1 and 4 after randomization or before discharge (if before day 4). Electrocardiograms and results of testing of cardiac biomarkers that were performed in hospitals for clinical indications were also collected. Samples were analyzed at the core laboratory of the Minneapolis Medical Research Foundation of Hennepin County Medical Center for troponin I (Access 2 Immunoassay System, Beckman Coulter) with the use of a threshold of 0.06  $\mu$ g per liter (1.5 times the 99th percentile [0.04  $\mu$ g per liter] for healthy patients). We used the Universal Definition of Myocardial Infarction criteria<sup>7,8</sup> to define myocardial infarction and unstable angina on the basis of review of clinical status, central interpretation of electrocardiograms at Saint Louis University, and results of core laboratory and clinical cardiac biomarkers (see the Supplementary Appendix, available at NEJM.org). Study investigators who classified cardiovascular outcomes and those who did follow-up telephone assessments were unaware of study-group assignments.

Other secondary outcomes that were determined on telephone follow-up at or close to 30 days and 60 days after randomization included current residence, survival, functional measures (lowerextremity physical and instrumental activities of daily living), and fatigue. These outcomes were ascertained with the use of methods described previously.<sup>6</sup>

#### TERTIARY STUDY OUTCOMES

We evaluated in-hospital morbidity up to 30 days after randomization, including pneumonia, wound infection, thromboembolism, stroke or transient ischemic attack, and clinically recognized myocardial infarction.<sup>6</sup> We prespecified two composite outcomes: death, myocardial infarction, or pneumonia; and death, myocardial infarction, pneumonia, thromboembolism, or stroke.

#### VITAL STATUS AND WALKING CONFIRMATION

We validated the vital status of patients in the United States by searching the online Social Security Database. When discrepancies were identified between telephone reports and this database, we verified deaths using hospital records or published obituaries. We validated the vital status of Canadian patients by searching hospital medical records, vital-status records, and outpatient medical records. We validated vital status in 95.9% of patients (99.0% in the United States and 91.2% in Canada). Of 1934 vital-status confirmations, we found 7 discrepancies (0.4%) between telephone re-

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ports and vital-status records; in these cases, we used vital-status records. We assessed the reliability of the self-report of walking status in a subgroup of 814 patients for whom we had both self-report and proxy report and found high reliability (kappa=0.90) between these reports.<sup>9</sup>

# ADHERENCE DEFINITIONS

We defined major protocol violations as a lack of receipt of a transfusion or hospital discharge with a hemoglobin level of less than 10 g per deciliter in the liberal-strategy group and as the receipt of transfusion with a hemoglobin level of 8 g per deciliter or more in the absence of symptoms in the restrictive-strategy group.

### STATISTICAL ANALYSIS

According to the original study design, we determined that a sample size of 2600 patients would provide a power of 90% and an experiment-wise alpha level of 0.05 allowing for interim analyses (four were performed by the data and safety monitoring board) and a level of 0.048 for the final comparison to detect an absolute between-group difference of 7 percentage points in the primary outcome (odds ratio, 0.75). In September 2007, the data and safety monitoring board approved a reduction of recruitment goal to 2000 patients. This change resulted in an absolute change of approximately 1 percentage point in the between-group difference in the primary outcome that could be excluded with a power of 90%.

We used the Mantel-Haenszel method<sup>10</sup> to conduct the primary analysis, taking into account different clinical sites. We prespecified tests for interaction of the primary outcome<sup>11,12</sup> with sex, age, race, and cardiovascular-disease status (known cardiovascular disease vs. risk factors only) without adjustment of the alpha level. Tests for interaction and differences in outcomes are presented without adjustment for clinical site. The primary outcome analysis is presented as a Mantel-Haenszel odds ratio with 95% confidence intervals. For secondary and tertiary analyses, we used standard methods for the comparison of proportions and means without adjustment for clinical site, using an alpha level of 0.01 (with 99% confidence intervals). Analyses were performed with the use of SAS software, version 9.2.

# RESULTS

# STUDY POPULATION

We screened 14,438 patients and randomly assigned 2016 to either the liberal-strategy group (1007 patients) or the restrictive-strategy group (1009) (see the Supplementary Appendix). There were 14 withdrawals, 2 losses to follow-up, and 1 incomplete follow-up ascertainment; follow-up for the primary analysis was obtained in 99.2% of the patients. Of the 1999 patients included in the primary analysis, we directly interviewed 1075 (53.8%) and obtained data on 923 (46.2%) by proxy; the source of information was missing for 1 patient.

The mean age of the study population was 81.6 years (range, 51 to 103), and cardiovascular disease was present in 62.9%. Baseline characteristics were similar in the two study groups (Table 1).

#### HEMOGLOBIN LEVELS AND TRANSFUSION

The average hemoglobin level before transfusion was 1.3 g per deciliter higher in the liberal-strategy group than in the restrictive-strategy group (P<0.001) (Table 2). The median number of units transfused was 2.0 (interquartile range, 1 to 2) in the liberal-strategy group and 0 (interquartile range, 0 to 1) in the restrictive-strategy group; 59.0% of patients in the restrictive-strategy group did not receive a transfusion after randomization. Figure 1 shows the average daily lowest hemoglobin levels in the two groups.

Violations in the transfusion protocol occurred in 9.0% of patients in the liberal-strategy group and in 5.6% of those in the restrictive-strategy group. Symptoms leading to transfusion are listed in Table 2.

#### OUTCOMES

The rates of death or an inability to walk without human assistance at 60-day follow-up were similar in the liberal-strategy group and the restrictive-strategy group (35.2% vs. 34.7%, P=0.90) (Table 3). The odds ratio for the primary outcome associated with the liberal strategy versus the restrictive strategy was 1.01 (95% confidence interval [CI], 0.84 to 1.22), for an absolute risk difference of 0.5 percentage points (95% CI, -3.7 to 4.7). There was a significant interaction according to patients' sex (P=0.03), with an odds ratio associ-

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Variable	Liberal Strategy (N = 1007)	Restrictive Strategy (N = 1009)
Age — yr	81.8±8.8	81.5±9.0
Male sex — no. (%)	250 (24.8)	239 (23.7)
Race — no. (%)†		
White	944 (93.7)	947 (93.9)
Black	40 (4.0)	42 (4.2)
Asian	14 (1.4)	13 (1.3)
Other	9 (0.9)	7 (0.7)
Residence in the United States — no. (%)	609 (60.5)	613 (60.8)
Cardiovascular disease — no. (%)		
Any	637 (63.3)	631 (62.5)
Coronary artery disease	402 (39.9)	403 (39.9)
Congestive heart failure	184 (18.3)	167 (16.6)
Cerebrovascular disease	249 (24.7)	224 (22.2)
Peripheral vascular disease	117 (11.6)	102 (10.1)
Cardiovascular risk factors — no./total no. (%)		
Hypertension	824/1003 (82.2)	821/1005 (81.7)
Diabetes mellitus	252/1003 (25.1)	256/1005 (25.5)
Hypercholesterolemia	347/1002 (34.6)	360/1001 (36.0)
Tobacco use	116/1003 (11.6)	113/1004 (11.3)
Creatinine >2.0 mg/dl	83/1001 (8.3)	86/1003 (8.6)
Chronic lung disease	189/1003 (18.8)	188/1007 (18.7)
History of dementia or confusion	309/1004 (30.8)	325/1008 (32.2)
History of cancer	181/1003 (18.0)	189/1008 (18.8)
Type of hip fracture — no./total no. (%)		
Femoral neck	432/1004 (43.0)	422/1008 (41.9)
Intertrochanteric	512/1004 (51.0)	522/1008 (51.8)
Subtrochanteric	88/1004 (8.8)	95/1008 (9.4)
Reverse oblique	13/1004 (1.3)	8/1008 (0.8)
Type of anesthesia — no./total no. (%)		
General	543/1005 (54.0)	566/1008 (56.2)
Spinal	457/1005 (45.5)	434/1008 (43.1)
Other	5/1005 (0.5)	8/1008 (0.8)
American Society of Anesthesiology risk score‡	3.0±0.6	2.9±0.6
Residence — no./total no. (%)		
Home or retirement home	892/1005 (88.8)	886/1008 (87.9)
Nursing home	104/1005 (10.3)	110/1008 (10.9)
Other	9/1005 (0.9)	12/1008 (1.2)

\* Plus-minus values are means ±SD. There were no significant between-group differences for any of the listed variables. † Race was self-reported.

‡ Scores range from 1 to 5, with a higher score indicating greater risk. Data in this category were missing for 38 patients in the liberal-strategy group and 39 in the restrictive-strategy group.

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Table 2. Hemoglobin Levels and Transfusions.*			
Variable	Liberal Strategy (N=1007)	Restrictive Strategy (N=1009)	P Value
Hemoglobin level — g/dl			
Before surgery	11.3±1.5	11.3±1.5	0.70
During eligibility screening	9.0±0.8	9.0±0.8	0.98
Before transfusion	9.2±0.5	7.9±0.6	<0.001
Estimated blood loss during surgery — ml $\dagger$	209±179	232±257	0.03
Transfusions before randomization			
0 units — no./total no. (%)	754/1006 (75.0)	720/1008 (71.4)	
≥1 unit — no./total no. (%)	252/1006 (25.0)	288/1008 (28.6)	0.07
Total no. of units	452	531	
Transfusions after randomization			
0 units — no./total no. (%)	33/1003 (3.3)	594/1007 (59.0)	
1 unit — no./total no. (%)	420/1003 (41.9)	246/1007 (24.4)	
2 units — no./total no. (%)	346/1003 (34.5)	127/1007 (12.6)	
3 units — no./total no. (%)	132/1003 (13.2)	24/1007 (2.4)	
≥4 units — no./total no. (%)	72/1003 (7.2)	16/1007 (1.6)	<0.001
Total no. of units	1866	652	
Storage of units transfused after randomization — days‡	22.0±9.5	22.1±9.9	0.83
Leukoreduced units transfused after randomization — $\%$	90.2	88.6	0.25
Major protocol violation — no./total no. (%) $\P$	91/1006 (9.0)	56/1007 (5.6)	0.003
Transfusion because of symptoms — no./total no. (%) $\ $			
Rapid bleeding	5/1006 (0.5)	14/1007 (1.4)	0.04
Chest pain	4/1006 (0.4)	9/1007 (0.9)	0.17
Congestive heart failure	1/1006 (0.1)	10/1007 (1.0)	0.007
Tachycardia or hypotension	43/1006 (4.3)	123/1007 (12.2)	<0.001

\* Plus-minus values are means ±SD.

† Data on estimated blood loss were missing for 122 patients in the liberal-strategy group and 129 in the restrictive-strategy group.

Data on the length of storage of units were missing for 25 units in the liberal-strategy group and 8 in the restrictivestrategy group.

§ Data on leukoreduction status were missing for 19 units in the liberal-strategy group and 10 in the restrictive-strategy group.

¶ In the liberal-strategy group, there were two types of protocol violations: 30 patients (3.0%) did not receive a transfusion, and 61 patients (6.1%) were discharged with a hemoglobin level of less than 10 g per deciliter. In the restrictivestrategy group, there was only one type of violation: 56 patients (5.6%) who did not have symptoms or rapid bleeding received transfusions for a hemoglobin level of 8.0 g per deciliter or more.

Patients may have had more than one symptom.

ated with the liberal strategy of 1.45 (95% CI, 1.00 to 2.10) for men versus 0.91 (95% CI, 0.74 to 1.13) for women. Interactions according to age, race, and cardiovascular-disease status were not significant (see the Supplementary Appendix).

There were no significant between-group differences in the rates of death on 30-day follow-up (5.2% in the liberal-strategy group vs. 4.3% in the restrictive-strategy group), for an absolute risk difference of 0.9 percentage points (99% CI, -1.5 to 3.4), and on 60-day follow-up (7.6% in the liberalstrategy group vs. 6.6% in the restrictive-strategy group), for an absolute risk difference of 1.0 percentage point (99% CI, -1.9 to 4.0) (Table 3). The between-group differences were also not significant in the rates of in-hospital acute myocardial infarction, unstable angina, or death (4.3% in the liberal-strategy group vs. 5.2% in the restrictive-

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strategy group), for an absolute risk difference of -0.9 percentage points (99% CI, -3.3 to 1.6). The frequencies of in-hospital clinical events and serious adverse events did not differ significantly between groups (Table 4). Also similar in the two groups were the length of hospital stay, scores for lower-extremity physical activities of daily living, instrumental activities of daily living, and fatigue, as well as rates of residing at home at 30-day and 60-day follow-up (Table 3).

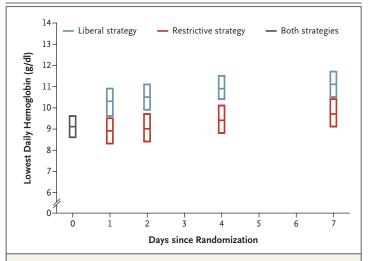
### DISCUSSION

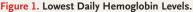
We performed a randomized clinical trial involving 2016 patients undergoing surgery for hip fracture and found no evidence that maintaining the hemoglobin level above 10 g per deciliter was superior to transfusion for symptoms or maintaining a hemoglobin level of less than 8 g per deciliter with respect to the primary outcome (a composite of death or an inability to walk across the room without human assistance) and to several clinically relevant secondary outcomes, including cardiovascular event rates and other functional measures. We enrolled a high-risk group of patients with a mean age of more than 81 years for whom untreated anemia would probably be more harmful than in a healthier or younger population undergoing most surgical procedures.

An ability to walk across the room at 60 days was selected as a main component of the primary outcome because such a measure is recognized to be an important functional outcome after hip fracture and is likely to be affected by factors that transfusion might influence (e.g., aerobic capacity and muscle strength). We hypothesized, in particular, that a higher hemoglobin level might facilitate more active participation in rehabilitation, leading to more successful recovery of ambulation.

We achieved a clinically important difference in the use of packed red cells and a good separation in hemoglobin levels in the two transfusion groups (Fig. 1). Patients in the restrictive-strategy group received 65% fewer units of blood than those in the liberal-strategy group; more than half the patients in the restrictive-strategy group did not receive any blood transfusion. Widespread implementation of this restrictive approach to transfusion in similar patients would greatly reduce blood use.

We found an interaction between the transfusion strategy and sex in the liberal-strategy group,





Shown are the lowest daily hemoglobin levels among patients in the liberalstrategy group versus those in the restrictive-strategy group. Data for the two groups are pooled on the day of randomization and are presented for days 1, 2, 4, and 7, when hemoglobin levels were required to be measured while patients remained in the hospital. The center line within each box represents the median, and the extremes the interquartile range.

suggesting a higher rate of death or an inability to walk without human assistance at 60-day follow-up in men but not in women. This difference was not anticipated and could have been due to chance.

We obtained primary-outcome information (including data regarding deaths) for more than 99% of patients and validated vital status. However, we did not perform follow-up examinations, and our telephone ascertainment of functional outcomes was subject to possible miscommunication, poorly informed proxy respondents, and recording errors. Although we did not validate patients' ability to walk, in cases in which both patients and their proxies answered the question about walking ability, we found strong agreement between the two reports. Scores for physical activities of daily living, instrumental activities of daily living, and fatigue were not validated and were not useful for analysis for 45 to 60% of patients. We revised eligibility criteria in the course of the trial to include lower-risk patients who had cardiovascular risk factors but no history of cardiovascular disease, and there was no important treatment interaction with cardiovasculardisease status.

Our study had excellent statistical power for determining the primary outcome of death or inability to walk. On the basis of the 95% confi-

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Table 3. Outcomes at 30 Days and 60 Days.*	Days.*							
Variable		30-Day Period	po			60-Day Period	riod	
	Liberal Strategy (N=1007)	Restrictive Strategy (N = 1009)	Odds Ratio (99% CI)	Absolute Risk Difference (99% CI)	Liberal Strategy (N=1007)	Restrictive Strategy (N = 1009)	Odds Ratio (95% CI)	Absolute Risk Difference (95% CI)
	no./tota	no./total no. (%)		percentage points	no./tota	no./total no. (%)		percentage points
Death or inability to walk indepen- dently	459/995 (46.1)	481/1000 (48.1)	0.92 (0.73 to 1.16)	-2.0 (-7.7 to 3.8)	351/998 (35.2)	347/1001 (34.7)	1.01 (0.84 to 1.22)	0.5 (-3.7 to 4.7)
Inability to walk independently	407/995 (40.9)	438/1000 (43.8)			275/998 (27.6)	281/1001 (28.1)		
Death	52/995 (5.2)	43/1000 (4.3)	1.23 0.9 (0.71 to 2.12) (-1.5 to 3.4)	0.9 (-1.5 to 3.4)	76/998 (7.6)	66/1001 (6.6)	1.17 1.0 (0.75 to 1.83)† (-1.9 to 4.0)†	1.0 (−1.9 to 4.0)†
			P Value	lue			P V	P Value
Residence			0.17	7			Ö	0.34
Home or retirement home	457/994 (46.0)	425/999 (42.5)			617/996 (61.9)	603/1001 (60.2)		
Nursing home	135/994 (13.6)	161/999 (16.1)			137/996 (13.8)	161/1001 (16.1)		
Other	402/994 (40.4)	413/999 (41.3)			242/996 (24.3)	237/1001 (23.7)		
	SCC	score			SC	score		
Function and symptom scales								
Lower-extremity physical ADL	7.3±4.0	7.4±3.9	0.72	2	5.1±4.2	$5.1 \pm 4.3$	.O	0.85
Instrumental ADL§	3.9±0.5	3.9±0.4	0.10	0	3.7±0.8	3.7±0.9	Ö	0.94
FACIT-Fatigue scale	38.7±7.7	38.6±7.6	0.84	4	41.8±7.3	<b>42.3</b> ±7.4	Ö	0.26
* Plus-minus values are means ±SD. Odds ratios and risk differences are for the comparison between the liberal-strategy group and the restrictive-strategy group.	odds ratios and risk	differences are for th	ie comparison b	etween the liber	al-strategy group a	nd the restrictive-strat	egy group.	

- Values are 99% confidence intervals.

number of dependencies with respect to 11 basic activities. Patients who reported that they had any human assistance in an activity or that they did not perform the activity for a health were excluded from the analysis. Scores were not used in this analysis on 30-day follow-up for 535 patients in the liberal-strategy group and 502 in the restrictive-strategy group and on Cores on the lower-extremity physical activities of daily living (ADL) scale range from 0 to 11, with higher scores indicating greater dependency. Scores were calculated by totaling the reason were considered to be dependent with respect to that activity. Patients who had missing data or who did not perform the activity for reasons other than those related to health 60-day follow-up for 484 in the liberal-strategy group and 456 in the restrictive-strategy group.

Scores on the instrumental ADL scale range from 0 to 4, with higher scores indicating greater dependency. Scores were calculated by totaling the number of dependencies with respect to four advanced activities. Patients who reported that they needed assistance or were unable to perform a task for health reasons were considered to be dependent with respect to that activity. Patients who had missing data or did not perform the activity for reasons other than those related to health were excluded from the analysis. Scores were not used in this analysis on 30-day follow-up for 570 patients in the liberal-strategy group and 559 in the restrictive-strategy group and on 60-day follow-up for 618 in the liberal-strategy group and 588 in the restrictive-strategy group.

The Functional Assessment of Chronic Illness Therapy-Fatigue (FACIT-Fatigue) scale includes 13 items with scores ranging from 0 to 4 (0, not at all; 1, a little bit; 2, somewhat; 3, quite a bit; and 4, very much), with higher scores indicating a greater energy level. Missing items were imputed as the mean of item scores within the same scale. No proxy responses were possible. Scores were missing on 30-day follow-up for 551 patients in the liberal-strategy group and 550 in the restrictive-strategy group and on 60-day follow-up for 463 in the liberalstrategy group and 484 in the restrictive-strategy group.

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Variable	Liberal Strategy (N=1007)	Restrictive Strategy (N=1009)	Odds Ratio (99% CI)	Absolute Risk Difference (99% CI)
	number/total i	number (percent)		percentage points
Myocardial infarction, unstable angina, or in-hospital death†	43/1005 (4.3)	52/1008 (5.2)	0.82 (0.48 to 1.42)	-0.9 (-3.3 to 1.6)
Myocardial infarction†	23/1005 (2.3)	38/1008 (3.8)	0.60 (0.30 to 1.19)	-1.5 (-3.5 to 0.5)
Unstable angina†	2/1005 (0.2)	3/1008 (0.3)	0.67 (0.06 to 7.03)	-0.1 (-0.7 to 0.5)
In-hospital death	20/1005 (2.0)	14/1008 (1.4)	1.44 (0.58 to 3.56)	0.6 (-0.9 to 2.1)
Isolated troponin elevation‡	62/1005 (6.2)	59/1008 (5.9)	1.06 (0.65 to 1.71)	0.3 (-2.4 to 3.1)
Physician diagnosis of congestive heart failure	27/1005 (2.7)	35/1007 (3.5)	0.77 (0.39 to 1.50)	-0.8 (-2.8 to 1.2)
Stroke or transient ischemic attack				
On CT or MRI	5/1005 (0.5)	1/1007 (0.1)	5.03 (0.30 to 84.73)	0.4 (-0.2 to 1.0)
On physician diagnosis or CT or MRI	8/1005 (0.8)	3/1007 (0.3)	2.69 (0.47 to 15.42)	0.5 (-0.3 to 1.3)
Chest radiograph with new or progressive infiltrate	60/1005 (6.0)	48/1007 (4.8)	1.27 (0.76 to 2.12)	1.2 (-1.4 to 3.8)
New-onset purulent sputum	9/1005 (0.9)	3/1007 (0.3)	3.02 (0.54 to 16.91)	0.6 (-0.3 to 1.5)
Wound infection	14/1005 (1.4)	8/1007 (0.8)	1.76 (0.56 to 5.56)	0.6 (-0.6 to 1.8)
Deep-vein thrombosis or pulmonary embolism	12/1005 (1.2)	8/1007 (0.8)	1.51 (0.46 to 4.92)	0.4 (-0.7 to 1.5)
Death, myocardial infarction, pneumonia	89/1005 (8.9)	90/1007 (8.9)	0.99 (0.66, 1.48)	-0.1 (-3.4 to 3.2)
Death, myocardial infarction, pneumonia, thrombo- embolism, or stroke	103/1005 (10.2)	94/1007 (9.3)	1.11 (0.75 to 1.63)	0.9 (-2.5 to 4.3)
Returned to operating room	15/1005 (1.5)	18/1007 (1.8)	0.83 (0.34 to 2.06)	-0.3 (-1.8 to 1.2)
Transfer to intensive care unit	30/1005 (3.0)	29/1007 (2.9)	1.04 (0.53 to 2.05)	0.1 (-1.8 to 2.0)
	days		PV	alue
Time from randomization to discharge∬				
United States	3.67±3.38	3.97±3.89	0.	15
Canada	12.03±9.31	12.70±9.48	0.	32

\* Plus-minus values are means ±SD. Odds ratios and risk differences are for the comparison between the liberal-strategy group and the restrictive-strategy group. CT denotes computed tomography, and MRI magnetic resonance imaging.

† Electrocardiographic results after randomization were incomplete for 135 patients in the liberal-strategy group and 130 in the restrictive-

strategy group. ‡ Blood samples obtained for troponin testing on day 4 after randomization or at the time of hospital discharge were not available for 180 patients in the liberal-strategy group and 175 in the restrictive-strategy group.

§ Of the 2011 patients who were evaluated (1220 in the United States and 791 in Canada), 944 patients (93.9%) in the liberal-strategy group and 934 (92.8%) in the restrictive-strategy group were discharged alive.

dence interval, the restrictive transfusion policy plausibly resulted in at most a 3.7% increase in the risk of death or inability to walk without human assistance, a composite outcome that occurred in about 35% of patients. We had less statistical power for in-hospital outcomes; our data are compatible with an absolute change in the composite outcome of in-hospital acute myocardial infarction, unstable angina, or death, ranging from an increase of 3.3 percentage points to a decrease of 1.6 percentage points for the restrictive transfusion strategy. Our results are consistent with most of the findings of the Transfusion Requirements in Critical Care (TRICC) trial, in which outcomes did not differ significantly between a transfusion threshold of 7 g per deciliter and a threshold of 10 g per deciliter among patients in intensive care units.<sup>5,13</sup> However, in contrast to that report, we did not find increased rates of myocardial infarction or congestive heart failure in the liberal-strategy group. Furthermore, we did not confirm findings from observational studies of markedly higher mortality in patients who received transfusion

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than in patients who did not.<sup>14</sup> A randomized clinical trial allows us to evaluate transfusion while avoiding selection bias.<sup>15</sup>

In summary, we found that a liberal transfusion strategy, as compared with a restrictive strategy, did not result in reduced rates of death or an inability to walk on 60-day follow-up or in significant reductions in rates of in-hospital complications in this population at increased cardiovascular risk. Our findings suggest that it is reasonable to withhold transfusion in patients who have undergone surgery in the absence of symptoms of anemia or a decline in the hemoglobin level below 8 g per deciliter, even in elderly patients with underlying cardiovascular disease or risk factors.

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Disclosure forms provided by the authors are available with the full text of this article at NEJM.org.

APPENDIX

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