

Review Article

The efficacy of lidocaine to prevent laryngospasm in children: a systematic review and meta-analysis*

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Summary

The purpose of this meta-analysis was to determine the efficacy of lidocaine in preventing laryngospasm during general anaesthesia in children. An electronic search of six databases was conducted. The Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) guidelines were adhered to. We included randomised controlled trials reporting the effects of intravenous and/or topical lidocaine on the incidence of laryngospasm during general anaesthesia. Nine studies including 787 patients were analysed. The combined results demonstrated that lidocaine is effective in preventing laryngospasm (risk ratio (RR) 0.39, 95% CI 0.24–0.66; $I^2 = 0$). Subgroup analysis revealed that both intravenous lidocaine (RR 0.34, 95% CI 0.14–0.82) and topical lidocaine (RR 0.42, 95% CI 0.22–0.80) lidocaine are effective in preventing laryngospasm. The results were not affected by studies with a high risk of bias. We conclude that, both topical and intravenous lidocaine are effective for preventing laryngospasm in children.

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Introduction

Laryngospasm is a serious complication of general anaesthesia that is known to occur most often at induction of anaesthesia, during tracheal intubation and at extubation. The reported frequency of laryngospasm during general anaesthesia is between 1.7% and 25% in children [1–3]. Laryngospasm causes complete obstruction of the upper airway and can lead to oxygen desaturation, negative pressure pulmonary oedema and death [4], and its prevention during general anaesthesia is challenging. Lidocaine, administered

either intravenously or topically, is often used to prevent laryngospasm in children during anaesthesia or immediately before tracheal extubation. However, studies have yielded conflicting results [5–8]. Most of the clinical studies are underpowered due to the relatively low incidence of laryngospasm during general anaesthesia. Moreover, the routes of administration (for example, intravenous or topical) have varied. We decided to analyse the efficacy of lidocaine in children and to identify the most effective route of administration. We therefore conducted a systematic review and

meta-analysis of randomised controlled trials (RCTs) that evaluated the effect of lidocaine in preventing laryngospasm during general anaesthesia.

Methods

We followed the recommendations of the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) statement [9, 10].

We searched MEDLINE, the Cochrane Central Register of Controlled Trials (CENTRAL), Embase and Web of Science initially on 15 February 2013 and again on 4 April 2014. The reference lists of the full articles were also searched. We also conducted a search of clinicaltrials.gov and the UMIN clinical trials registry. We used the following search strategy combining free text and MeSH terms for searching PubMed: (“laryngismus”[MeSH Terms] OR “laryngismus”[All Fields] OR “laryngospasm”[All Fields] OR “laryngeal spasm”[All Fields] OR “laryngeal spasms”[All Fields] OR “emergence”[All Fields]) AND (“lidocaine”[MeSH Terms] OR “lidocaine”[All Fields] OR “lignocaine”[All Fields] OR “xylocaine”[All Fields]) AND (“controlled clinical trial”[pt] OR “randomized”[tiab] OR “placebo”[tiab] OR “drug therapy”[Subheading] OR “randomly”[tiab] OR “trial”[tiab] OR “groups”[tiab]) NOT (“animals”[MeSH Terms] NOT “humans”[MeSH Terms]). Two authors (TM and KU) independently assessed the suitability of all the studies that were identified for potential inclusion as a result of the search strategy. The full-text versions of potentially relevant studies chosen by at least one author were retrieved and evaluated and those that met the inclusion criteria were then assessed separately. Any discrepancies were resolved through discussion and if no agreement could be reached, a third author (TG) arbitrated.

All trials that compared lidocaine with a control (i.e. placebo or no treatment) and reported the incidence of laryngospasm were included in this study. We excluded studies in adults and studies that did not have adequate comparators. We excluded studies in which we could not confirm the incidence of laryngospasm because it was not clearly reported. We also excluded case reports, comments or letters to the editor, reviews and animal studies. Eligibility was not restricted by language, type of surgery or anaesthetic technique.

The primary outcome of the meta-analysis was evaluation of the effect of lidocaine, compared with a control, for prevention of laryngospasm during general anaesthesia in children. We conducted subgroup analysis according to the route of administration (i.e. intravenous or topical lidocaine).

A data collection sheet was created that included information on: (1) ASA physical status; (2) age; (3) type of surgery; (4) anaesthetic technique; (5) type of airway device; (6) route of administration of lidocaine; (7) dose of lidocaine; (8) timing of administration of lidocaine; (9) timing of assessment for laryngospasm; (10) type of control (i.e. placebo or no treatment); (11) primary outcomes of individual studies; number of patients in (12) lidocaine group and (13) control group; number of reported incidents of laryngospasm in (14) lidocaine group and (15) control group; and (16) adverse effects of lidocaine such as seizures, arrhythmias or allergic reactions. When laryngospasm was classified according to severity (e.g. mild, moderate or severe), we extracted the data from the severe category; we did not consider stridor or bronchospasm as laryngospasm. If laryngospasm was reported together with stridor or bronchospasm and we could not extract the data of laryngospasm separately, we contacted the lead author for more information. If we were unable to obtain more detailed data from the author, we extracted the mixed data, i.e. laryngospasm, stridor and bronchospasm. We conducted a sensitivity analysis to confirm whether the pooled results would change when the mixed data were excluded.

Two authors (TM and KU) independently extracted the data from the included studies and then cross-checked the data. Disagreements were resolved by discussion between the two authors. If no agreement could be reached, a third author (TG) arbitrated.

A translator was consulted for studies published in languages other than English or Japanese.

We assessed the risk of bias as described by the Cochrane Handbook for Systematic Reviews of Interventions [11]. We assessed the risk of bias in sequence generation, allocation sequence concealment, blinding of patients, blinding of healthcare providers, blinding of data collectors, blinding of outcome

assessors, incomplete outcome data, selective outcome reporting and other bias. We also summarised risk of bias. Because we determined that a lack of blinding of patients would be unlikely to affect the incidence of laryngospasm, we attributed a low risk of bias to all studies in this domain. Two authors (TM and KU) independently assessed risk of bias for each RCT. Disagreements were resolved by a third author (TG).

For evaluating the quality of evidence for the effect of lidocaine in preventing laryngospasm, we used the Grading of Recommendations Assessment, Development and Evaluation (GRADE) approach [12, 13], which specifies four levels: high; moderate; low; and very low.

Dichotomous data were summarised using risk ratio (RR) with a 95% CI. If the 95% CI included a value of 1, we considered the difference to be non-significant. The number needed to treat (NNT) was calculated to estimate the overall clinical impact of the intervention. Heterogeneity was quantified with the I^2 statistic. We used the random-effects model (Dersimian and Laird method) to combine the results of the studies. Forest plots were used to represent and evaluate the effects of treatment graphically. Small study effects, including publication bias, were assessed using a contour-enhanced funnel plot and a Begg's rank correlation test [14] and was considered to be present at a p value < 0.1 in the asymmetry test. Subgroup analyses were performed according to the route of administration (intravenous or topical). Sensitivity analyses were performed by excluding studies with a high risk of bias; restricting to peer reviewed manuscripts; and excluding studies that reported mixed data for the occurrence of laryngospasm, stridor and bronchospasm. In addition, we conducted meta-regression analysis incorporating the following covariates: type of surgery (tonsillectomy vs others); airway device (tracheal tube vs supraglottic airway device); definition of 'laryngospasm' (mixed data vs others); route of administration (topical vs intravenous); and dose of lidocaine (low vs high). We defined an intravenous dose $> 1.5 \text{ mg.kg}^{-1}$ or a topical dose of 4 mg.kg^{-1} as 'high-dose'. Statistical analyses were performed using the R statistical software package version 2.13.0 (R Foundation for Statistical Computing, Vienna, Austria).

Results

Our search of MEDLINE, the Cochrane Central Register of Controlled Trials (CENTRAL), Embase, Web of Science, clinicaltrials.gov and the UMIN clinical trial registry databases produced a total of 535 citations. The full texts of 49 articles [5–8, 15–59] were examined in detail. We included nine RCTs [5–7, 38, 44–48] with a total of 787 participants (Fig. 1). Of the included studies, eight were available in English [5–7, 38, 45–48] and one in Spanish [44]. Eight studies were presented as peer-reviewed papers [5–7, 44–48] and one as an unpublished abstract [38].

The details of all included studies are shown in Table 1. Lidocaine was administered intravenously in five studies [5–7, 38, 47] and topically in five studies [6, 44–46, 48]. Intravenous lidocaine was administered in doses between 1.0 and 2.0 mg.kg^{-1} before tracheal extubation. In four of these studies, intravenous lidocaine was administered within 5 min of extubation [5–7, 47], while the timing of administration was described simply as 'before extubation' in another study [38].

Methods of topical lidocaine application included lidocaine spray on to the larynx [6], lidocaine spray to the supraglottic, glottic and subglottic areas [48], aerosol administration [44] or lidocaine jelly placed on the dorsal surface of the supraglottic airway device [45, 46]. Topical lidocaine was administered before induction in one study [44], before intubation in one study [6], at the time of intubation in one study [48] and during supraglottic airway device insertion in two studies [45, 46].

There were three studies that reported mixed data for laryngospasm, stridor and bronchospasm [6, 46, 48]. Adverse effects of lidocaine were not reported in these studies.

The combined results from nine studies that included 787 patients demonstrated a statistically significant reduction in the incidence of laryngospasm in patients who received lidocaine. In addition, subgroup analysis indicated that both intravenous and topical lidocaine have a statistically significant effect on reducing the incidence of laryngospasm (Fig. 2). The reported frequency of laryngospasm during general anaesthesia in children is between 1.7% and 25%

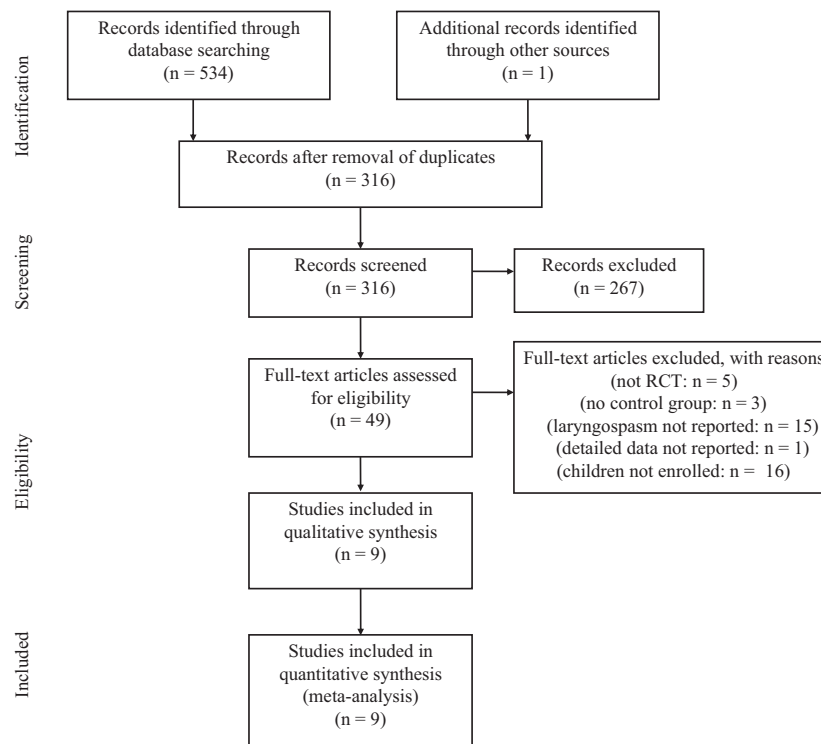


Figure 1 PRISMA flow diagram showing the identification, screening, eligibility, and inclusion of studies on the effect of lidocaine in preventing laryngospasm in children. RCT, randomised controlled trial.

[1–3]. The number needed to treat (NNT) to prevent laryngospasm in children is between 7 (95% CI 5–12) and 96 (95% CI 77–173) when considering these baseline risks.

The results were not different when studies with a high risk of bias were excluded (RR 0.44, 95% CI, 0.24–0.82), when restricting data to peer-reviewed manuscripts (RR 0.41, 95% CI 0.24–0.69) or when studies reporting mixed data for laryngospasm were excluded RR (0.31, 95% CI 0.12–0.82). The results of meta-regression analysis indicated that none of these five covariates had a significant effect (Table 2). A low risk of bias was attributed to one study, and a high risk of bias was attributed to four studies (Table 3). All nine studies mentioned randomisation, but only two studies reported the details of randomisation and only one study reported details of the method of allocation concealment. Healthcare providers, data collectors and outcome assessors were adequately blinded in five, three and three studies, respectively.

According to the GRADE methodology, the quality of the evidence for the effect of lidocaine in

preventing laryngospasm was downgraded to moderate because of a serious risk of bias.

Small study effects were assessed using contour-enhanced funnel plot (Fig. 3) and Begg's rank correlation test. The results of the asymmetry test were not statistically significant.

Discussion

The results of our meta-analysis demonstrate that lidocaine is effective in preventing laryngospasm in children during general anaesthesia. Our sensitivity analyses have revealed that the effect of lidocaine in preventing laryngospasm in children is not affected by studies with a high risk of bias and/or non-peer-reviewed studies. The quality of evidence, according to the GRADE approach, was moderate. In addition, our subgroup analyses indicate that both intravenous and topical lidocaine are effective in preventing laryngospasm in children.

Our systematic review and meta-analysis reveal that the risk ratio (95% CI) of laryngospasm during general anaesthesia in children is 0.39 (0.24–0.66).

Table 1 A summary of the included studies that assessed the effect of lidocaine in preventing laryngospasm in children.

Source	ASA physical status	Age	Surgery	Anaesthetic technique	Type of airway device	Route of administration	Dose of lidocaine	Timing of administration	Control
Sanikop & Bhat 2010 [47]	1, 2	3 months – 6 years	Cleft palate surgery	Ketamine and suxamethonium (induction), N ₂ O and vecuronium (maintenance)	Tracheal tube	Intravenous	1.5 mg.kg ⁻¹	2 min before extubation	Placebo
Koç et al. 1998 [6]	Not mentioned	5–10 years	Tonsillectomy and adenoidectomy	N ₂ O and halothane	Tracheal tube	Intravenous	1 mg.kg ⁻¹	5 min before extubation	Placebo
Leicht et al. 1985 [7]	1	3–7 years	Tonsillectomy	N ₂ O and halothane (induction), N ₂ O and halothane (maintenance)	Tracheal tube	Intravenous	1.5 mg.kg ⁻¹	4–5 min before extubation	Placebo
Bidwai et al. 1979 [38]	1	2–8 years	Tonsillectomy and adenoidectomy	N ₂ O and halothane	Tracheal tube	Intravenous	1 mg.kg ⁻¹	Before extubation	Placebo
Baraka 1978 [5]	Not mentioned	3–6 years	Tonsillectomy	Halothane (induction), halothane (maintenance)	Tracheal tube	Intravenous	2 mg.kg ⁻¹	One minute before extubation	No treatment
Schebesta et al. 2010 [46]	1, 2	1–10 years	Minor surgical procedures	Sevoflurane, fentanyl, and propofol (induction), fentanyl and sevoflurane (maintenance)	Supraglottic airway device (SAD)	Topical	0.3 ml.kg ⁻¹ of 2% lidocaine-containing gel	During insertion of SAD	Placebo
Penaloza et al. 1999 [44]	1, 2	1–8 years	Not mentioned	Halothane	Tracheal tube	Topical	10 mg	Before induction	Placebo
Koç et al. 1998 [6]	Not mentioned	5–9 years	Tonsillectomy and adenoidectomy	N ₂ O and halothane	Tracheal tube	Topical	4 mg.kg ⁻¹	Before intubation	Placebo
O'Neill et al. 1994 [45]	Not mentioned	4 months – 14 years	Not mentioned	N ₂ O and halothane (induction), N ₂ O and halothane or isoflurane (maintenance)	Supraglottic airway device	Topical	Approximately 1/4 teaspoon of 2% viscous lidocaine	During insertion of SAD	Placebo
Staffel et al. 1991 [48]	Not mentioned	10 years	Tonsillectomy and/or adenoidectomy	Not mentioned	Tracheal tube	Topical	4 mg.kg ⁻¹	At the time of intubation	No treatment

N₂O, nitrous oxide.

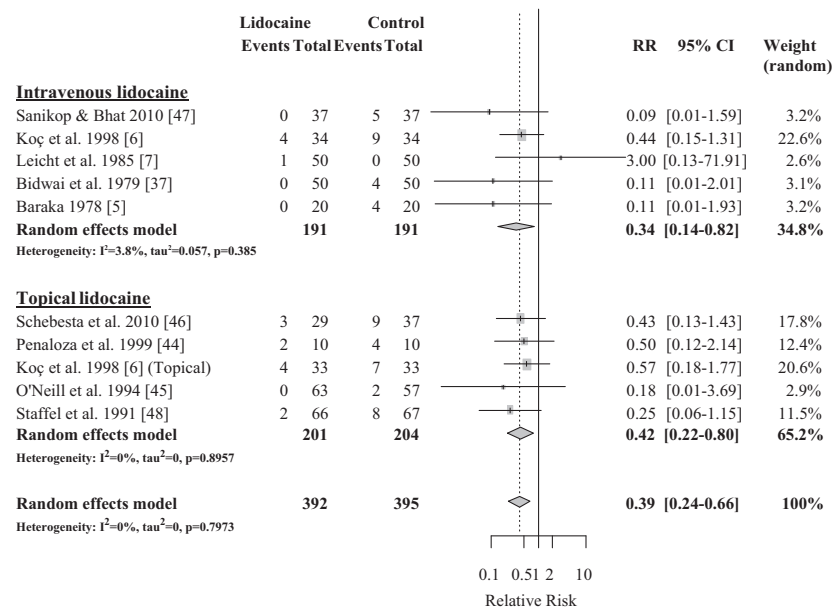


Figure 2 Forest plot for subgroup analysis of the effect of lidocaine on preventing laryngospasm in children.

Table 2 Results of a meta-regression analysis to determine whether covariates had a significant effect.

	Relative change in risk ratio (95% CI)	p value
Type of surgery	1.17 (0.09–15.0)	NS
Airway device	1.43 (0.16–13.1)	NS
Definition of laryngospasm	0.74 (0.12–4.37)	NS
Route of administration	1.55 (0.31–7.75)	NS
Dose of lidocaine	1.44 (0.31–6.74)	NS

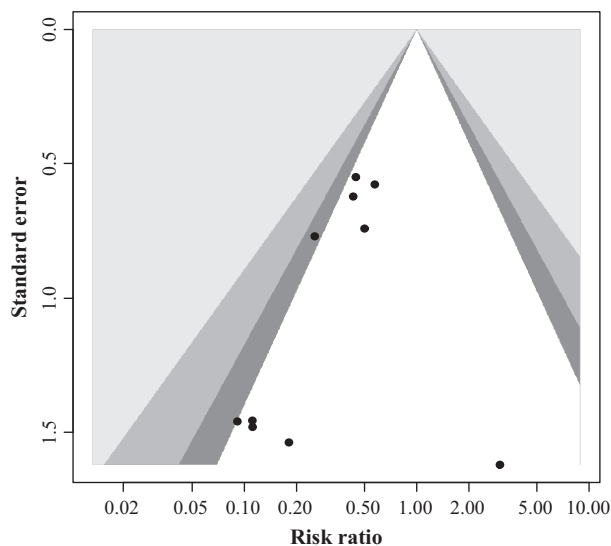
Most RCTs could not provide good estimates of the effect because the incidence of laryngospasm is low. Recently, however, Erb et al. conducted a crossover trial [60] to investigate the effect of intravenous lidocaine in preventing laryngospasm. In their study, laryngospasm was induced by spraying distilled water onto the larynx in children anaesthetised with sevoflurane and hence the incidence of laryngospasm was increased. Their study design allowed them to provide better estimates of the effect of lidocaine. They reported that the incidence of laryngospasm was reduced from 38% at baseline to 15% 2 min after 2 mg.kg⁻¹ intravenous lidocaine administration; the risk ratio was 0.4, which is almost identical to ours and, we believe, confirms that our results provide the best estimate thus far of the efficacy of lidocaine.

Our meta-analysis included four studies [5–7, 47] in which intravenous lidocaine was administered within 5 min of tracheal extubation and, in another study [38], the timing of administration was described as 'before extubation'. It has been reported [35] that coughing after extubation can be prevented by 2 mg.kg⁻¹ intravenous lidocaine administered within 5 min of tracheal extubation, but this effect was short-lived because the plasma concentration of lidocaine was only sufficiently high (> 3 µg.ml⁻¹) for 5 min after intravenous administration. In the crossover trial by Erb et al. [60], the incidence of laryngospasm 2 min following 2 mg.kg⁻¹ intravenous lidocaine administration was reduced significantly, while the effect was no longer significant at 10 min. The timing of intravenous administration of lidocaine appears to be important in its efficacy in preventing laryngospasm. In our meta-analysis, no study evaluating the effect of intravenous lidocaine administered > 5 min before extubation was included. Thus, we believe that prophylactic intravenous lidocaine for the prevention of laryngospasm should be administered within 5 min of tracheal extubation.

A recent upper respiratory tract infection (URTI) and [47] passive smoking are known to increase the risk of laryngospasm in children. Schebesta et al. [46]

Table 3 Risk of bias in studies assessing the effect of lidocaine in preventing laryngospasm in children.

Source	Sequence generation	Allocation concealment	Patients blinded	Healthcare providers blinded	Data collectors blinded	Outcome assessors blinded	Incomplete outcome data	Selective reporting	Other bias	Summary
Sanikop & Bhat 2010 [47]	Low	Unclear	Low	Unclear	Unclear	Unclear	Unclear	Low	Low	Unclear
Koç et al. 1998 [6] (Intravenous)	Unclear	Unclear	Low	Low	Unclear	Unclear	Unclear	Low	Low	Unclear
Leicht et al. 1985 [7]	Unclear	Unclear	Low	Low	Low	Low	Unclear	Low	Low	Unclear
Bidwai et al. 1979 [38]	Unclear	Unclear	Low	Low	Low	Low	Unclear	Unclear	Unclear	Unclear
Baraka 1978 [5]	Unclear	Unclear	Low	High	High	High	Unclear	Unclear	Unclear	High
Schebesta et al. 2010 [46]	Low	Low	Low	Low	Low	Low	Low	Low	Low	Low
Penaloza et al. 1999 [44]	Unclear	Unclear	Low	High	High	High	Unclear	Unclear	Unclear	High
Koç et al. 1998 [6] (Topical)	Unclear	Unclear	Low	Low	Unclear	Unclear	Unclear	Low	Low	Unclear
O'Neill et al. 1994 [45]	Unclear	Unclear	Low	High	High	High	Low	Low	Low	High
Staffel et al. 1991 [48]	High	High	Low	High	High	High	Unclear	Low	Low	High

**Figure 3** Contour-enhanced funnel plot to assess small study effects of studies included in the meta-analysis. (■ 0.1 > p > 0.05, ▒ 0.05 > p > 0.01, □ p < 0.01, • studies)

reported that topical lidocaine decreased the incidence of laryngospasm in children who had suffered a recent URTI. This was the only study we found that investigated the effect of lidocaine in preventing laryngospasm during general anaesthesia in these high-risk

patients and we suggest that the efficacy of lidocaine in high-risk children, such as those exposed to passive smoking, should be investigated.

An important limitation to our meta-analysis concerned risk of bias. Anaesthetists were not blinded in four studies, and most studies did not describe the randomisation process or comment on allocation concealment. As a result, a high risk of bias was attributed to four studies and a low risk of bias was attributed to only one study. Thus, we downgraded the quality of evidence to moderate. Further studies with a more robust design are required to confirm our findings. In spite of this limitation, our systematic review and meta-analysis has important strengths. Firstly, the extensive search strategy without language restriction using different databases, including a pre-registration site, allowed us to retrieve many studies. As a result, we included one article in Spanish [44] and one published in abstract form only [38], which may contribute to reduce publication bias.

In conclusion, the results of this meta-analysis demonstrate that lidocaine is able to prevent laryngospasm during general anaesthesia in children (GRADE: moderate) and that both intravenous and topical administration are effective.

Competing interests

No external funding and no competing interests declared.

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